



WESTFÄLISCHE  
WILHELMS-UNIVERSITÄT  
MÜNSTER

## MCMC Sampling for Bayesian Inference using L1-type Priors

*(what I do whenever the ill-posedness of EEG/MEG is just not frustrating enough!)*

AG Imaging Seminar

## Sparsity Constraints in Inverse Problems

Current trend in high dimensional inverse problems: **Sparsity constraints**.

- ▶ **Total Variation (TV)** imaging: Sparsity constraints on the gradient of the unknowns.
- ▶ **Compressed Sensing**: High quality reconstructions from a small amount of data, if a sparse basis/dictionary is a-priori known (e.g., wavelets).

Some nice images here!

## Sparsity Constraints in Inverse Problems

Commonly applied formulation and analysis by means of **variational regularization**, mostly by incorporating L1-type norms:

$$\hat{u}_\alpha = \operatorname{argmin}_{u \in \mathbb{R}^n} \left\{ \|m - Au\|_2^2 + \alpha \|Du\|_1 \right\}$$

assuming additive Gaussian i.i.d. noise  $\sim \mathcal{N}(0, \sigma^2)$

Notation:

- ▶  $m \in \mathbb{R}^k$ : The noisy measurement data given
- ▶  $u \in \mathbb{R}^n$ : The unknowns to recover w.r.t. the chosen discretization
- ▶  $A \in \mathbb{R}^{k \times n}$ : Discretization of the forward operator w.r.t. the domains of  $u$  and  $m$ .
- ▶  $D \in \mathbb{R}^{l \times n}$ : Discrete formulation of the mapping onto the (potentially) sparse quantity.

## Sparsity Constraints in Inverse Problems

Sparsity constraints relying on L1-type norms can also be formulated in the Bayesian framework.

- ▶ **Likelihood** model:

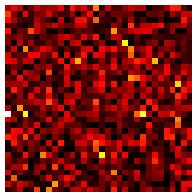
$$M = A u + \mathcal{E} \quad \mathcal{E} \sim \mathcal{N}(0, \sigma^2 I_k) \quad p_{li}(m|u) \propto \exp\left(-\frac{1}{2\sigma^2} \|m - A u\|_2^2\right)$$

- ▶ **Prior** model:

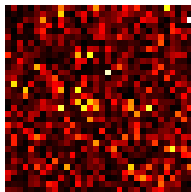
$$p_{pr}(u) \propto \exp(-\lambda |D u|_1)$$

- ▶ Resulting **posterior**:

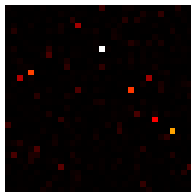
$$p_{post}(u|m) \propto \exp\left(-\frac{1}{2\sigma^2} \|m - A u\|_2^2 - \lambda |D u|_1\right)$$



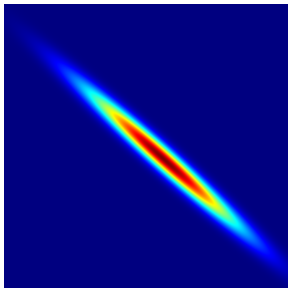
(a)  $\exp\left(-\frac{1}{2} \|u\|_2^2\right)$



(b)  $\exp(-|u|_1)$

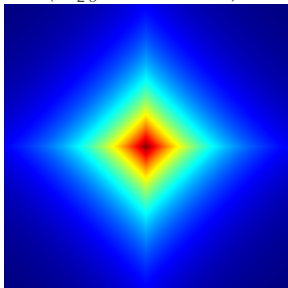


(c)  $1/(1 + u^2)$



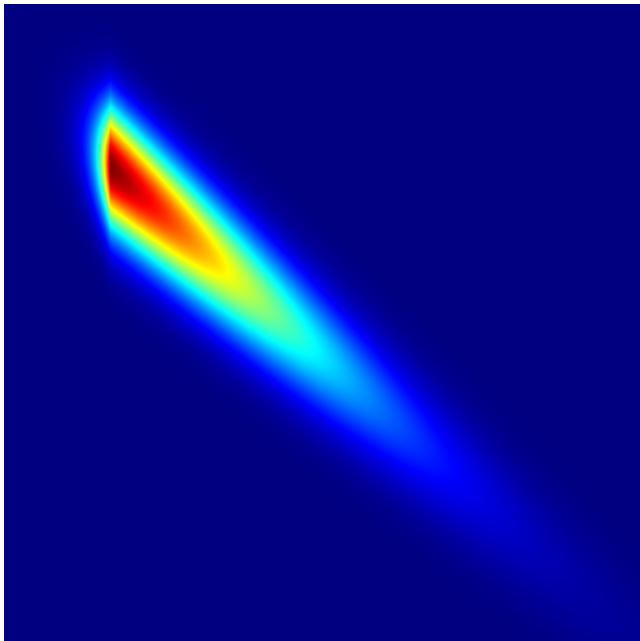
Likelihood:

$$\exp\left(-\frac{1}{2\sigma^2}\|m - Au\|_2^2\right)$$



Prior:  $\exp(-\lambda |u|_1)$

( $\lambda$  via discrepancy principle)



Posterior:  $\exp\left(-\frac{1}{2\sigma^2}\|m - Au\|_2^2 - \lambda |u|_1\right)$

## Sparsity Constraints in Inverse Problems

Direct correspondence to variational regularization by **maximum a-posteriori-estimation (MAP)** inference strategy:

$$\begin{aligned}\hat{u}_{\text{MAP}} &:= \operatorname{argmax}_{u \in \mathbb{R}^n} p_{\text{post}}(u|m) \\ &= \operatorname{argmax}_{u \in \mathbb{R}^n} \left\{ \exp \left( -\frac{1}{2\sigma^2} \|m - Au\|_2^2 - \lambda |Du|_1 \right) \right\} \\ &= \operatorname{argmin}_{u \in \mathbb{R}^n} \left\{ \|m - Au\|_2^2 + 2\sigma^2\lambda |Du|_1 \right\}\end{aligned}$$




⇒ Properties of MAP estimate (e.g., *discretization invariance*) are well understood.

## Sparsity Constraints in Inverse Problems

But there is more to Bayesian inference:

- ▶ Conditional mean-estimates (CM)
- ▶ Confidence intervals estimates
- ▶ Conditional covariance estimates
- ▶ Histogram estimates
- ▶ Generalized Bayes estimators
- ▶ Marginalization
- ▶ Model selection or averaging
- ▶ Experiment design

Influence of sparsity constraints on these quantities: **Less well understood.**

-  M. Lassas and S. Siltanen, 2004.  
Can one use total variation prior for edge-preserving Bayesian inversion?
-  M. Lassas, E. Saksman, and S. Siltanen, 2009.  
Discretization invariant Bayesian inversion and Besov space priors.
-  V. Kolehmainen, M. Lassas, K. Niinimäki, and S. Siltanen, 2012.  
Sparsity-promoting Bayesian inversion.

## Sparsity Constraints in Inverse Problems

### Key issue

Examining L1-type priors might help to further understand the relation between variational regularization theory and Bayesian inference!

### Key problem

Bayesian inference relies on computing integrals w.r.t. high-dim. posterior  $p_{post}$ . Standard Monte Carlo integration techniques break down for ill-posed, high-dimensional problems with sparsity constraints.

This talks summarizes partial results from:



F. Lucka, 2012.

Fast MCMC sampling for sparse Bayesian inference in high-dimensional inverse problems using L1-type priors

*submitted to Inverse Problems; arXiv:1206.0262v1*



## Monte Carlo Integration in a Nutshell

$$\mathbb{E}[f(x)] = \int_{\mathbb{R}^n} f(x) p(x) dx$$

- ▶ *Traditional Gauss-type quadrature:*

Construct suitable grid  $\{x_i\}_i$ , w.r.t  $\omega(x) := p(x)$  and approximate by  $\sum_{i=1}^K \omega_i f(x_i)$ .  
 $\implies$  Grid construction and evaluation **infeasible** in high dimensions.

- ▶ *Monte Carlo integration idea:*

Generate suitable grid  $\{x_i\}_i$ , w.r.t  $p(x)$  by drawing  $x_i \sim p(x)$  and approximate by  $\frac{1}{K} \sum_{i=1}^K f(x_i)$ . By the *Law of large numbers*:

$$\frac{1}{K} \sum_{i=1}^K f(x_i) \xrightarrow{K \rightarrow \infty} \mathbb{E}_{p(x)}[f(x)] = \int_{\mathbb{R}^n} f(x) p(x) dx$$

in L1 with rate  $O(K^{-1/2})$  (**independent of  $n$** ).

## Markov Chain Monte Carlo

Not able to draw **independent** samples?

↪ With  $\{x_i\}_i$  being an **ergodic Markov chain**, it still works!

**Markov chain Monte Carlo** (MCMC) methods are algorithms to construct such a chain:

- ▶ Huge number of MCMC methods exists.
- ▶ No “universal” method.
- ▶ Most methods rely on one two basic schemes:
  - ▶ Metropolis-Hastings (MH) Sampling [Metropolis et al., 1953; Hastings, 1970]
  - ▶ Gibbs Sampling [Geman & Geman, 1984]
- ▶ Posteriors from inverse problems seem to be “special”.

In this talk: Comparison between the **most basic variants** of MH and Gibbs sampling for high-dimensional posteriors from inverse problem scenarios.

## Symmetric, Random-Walk Metropolis-Hastings Sampling

Given: Density  $p(x)$ ,  $x \in \mathbb{R}^n$  to sample from.

Let  $p_{pro}(z)$  be a symmetric density in  $\mathbb{R}^n$  and  $x_0 \in \mathbb{R}^n$  an initial state. Define burn-in size  $K_0$  and sample size  $K$ .

For  $i = 1, \dots, K_0 + K$  do:

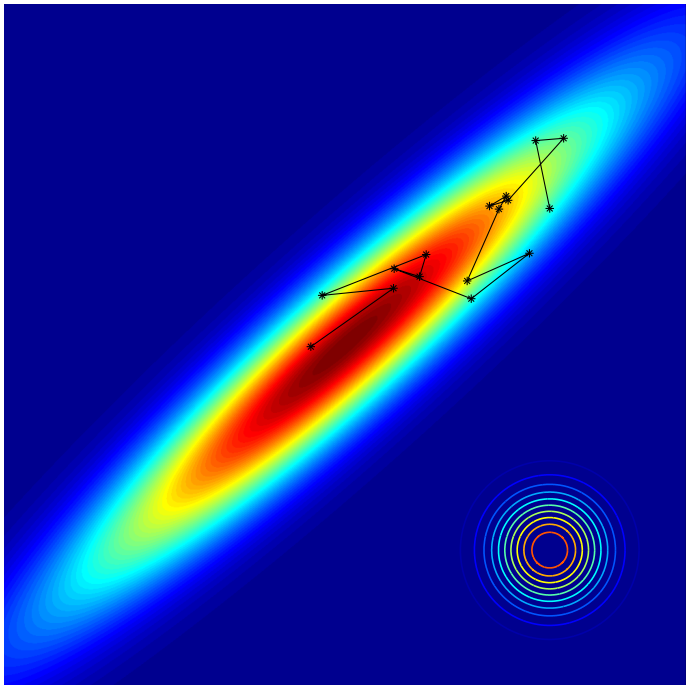
- 1 Draw  $z$  from  $p_{pro}(z)$  and set  $y = x_{i-1} + z$
- 2 Compute the acceptance ratio  $r = \frac{p(y)}{p(x_{i-1})}$
- 3 Draw  $\theta \in [0, 1]$  from a uniform probability density.
- 4 If  $r \geq \theta$ , set  $x_i = y$ , else set  $x_i = x_{i-1}$ .

Return  $x_{K_0+1}, \dots, x_K$ .

- ▶ Requires one evaluation of  $p(x)$  and one sample from  $p_{pro}$  per step, no “real” knowledge about  $p$  is needed, not even normalization.  
     $\rightsquigarrow$  “Black box” sampling algorithm.
- ▶ Most widely used.
- ▶ Good performance requires careful tuning of  $p_{pro}$ .
- ▶ Basis for very sophisticated sampling algorithms.
- ▶ Simulated annealing for the global optimization works in the same way.

In this talk:

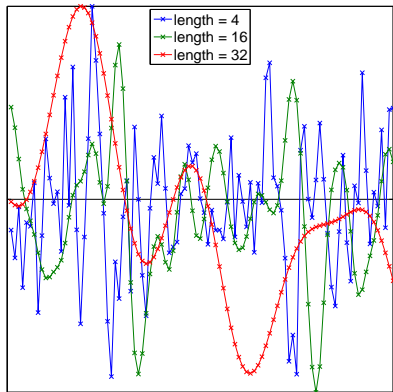
$$p_{pro} = \mathcal{N}(0, \kappa^2 I_n)$$



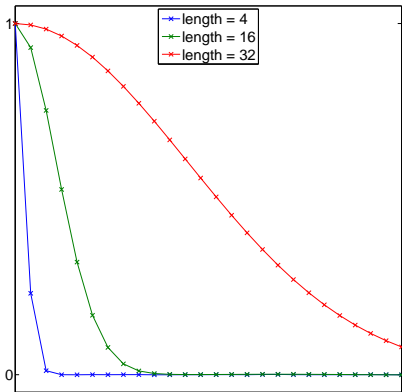
Evaluate performance of a sampler via its **autocorrelation function (acf)**:

Let  $g : \mathbb{R}^n \rightarrow \mathbb{R}^1$ , and  $g_i := g(u_i)$ ,  $i = 1, \dots, K$ , then

$$R(\tau) := \frac{1}{(K - \tau)\hat{\sigma}} \sum_{i=1}^{K-\tau} (g_i - \hat{\mu})(g_{i+\tau} - \hat{\mu}) \quad (\text{"lag-}\tau \text{ ac w.r.t. } g\text{"})$$

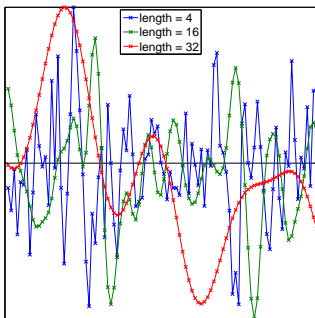


(a) Stochastic processes...

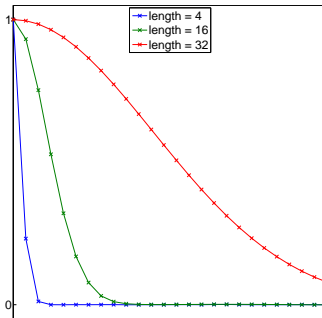


(b) ...and their autocorrelation functions

- ▶ A rapid decay of  $R(\tau)$  means that samples get uncorrelated fast.
- ▶ Temporal acf (tacf): acf rescaled by computation time per sample,  $t_s$ :  
 $R^*(t) := R(t/t_s)$  for all  $t = i \cdot t_s, i \in \{0, \dots, K - 1\}$ .
- ▶ Use  $g(u) := \langle \nu_1, u \rangle$ , where  $\nu_1$  is the largest eigenvector of the covariance matrix of  $p(x)$  to test the “worst case”.



(c) Stochastic processes...

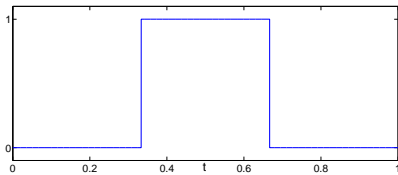


(d) ...and their autocorrelation functions

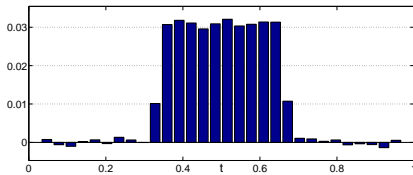
## Total Variation Deblurring Example in 1D (from Lassas & Siltanen, 2004)

- ▶ Model of a **charge coupled device** (CCD) in 1D.
- ▶ Unknown light intensity  $\tilde{u} : [0, 1] \rightarrow \mathbb{R}^+$ , indicator on  $[\frac{1}{3}, \frac{2}{3}]$ .
- ▶ Integrated into  $k = 30$  CCD pixels  $[\frac{1}{k+2}, \frac{k+1}{k+2}] \subset [0, 1]$ .
- ▶ Noise is added.
- ▶  $\tilde{u}$  is reconstructed on a regular,  $n$ -dim. grid.
- ▶  $D$  is the forward finite difference operator with NB cond.

$$p_{\text{post}}(u|m) \propto \exp\left(-\frac{1}{2\sigma^2}\|m - Au\|_2^2 - \lambda |Du|_1\right)$$



(e) The unknown function  $\tilde{u}(t)$



(f) The measurement data  $m$

## Total Variation Deblurring Example in 1D (from Lasso & Siltanen, 2004)

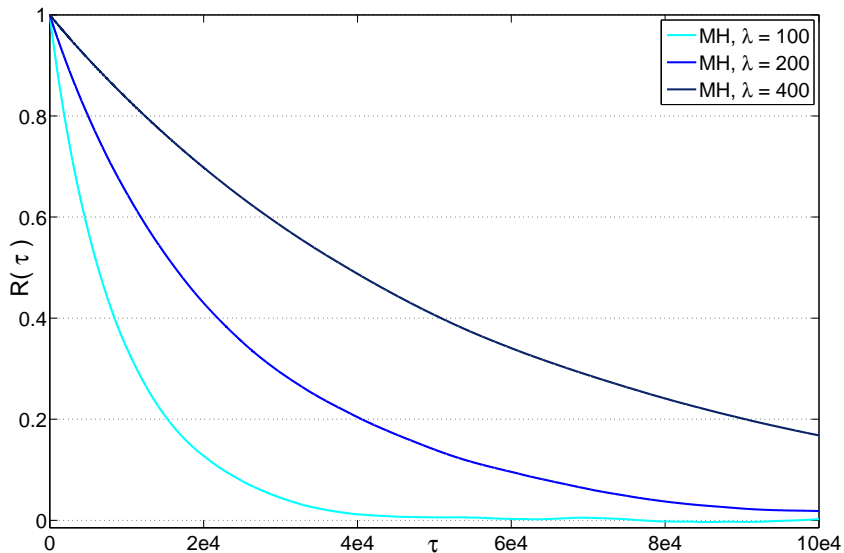


Figure: Autocorrelation plots  $R(\tau)$  for MH Sampler and  $n = 63$ .



## Total Variation Deblurring Example in 1D (from Lasso & Siltanen, 2004)

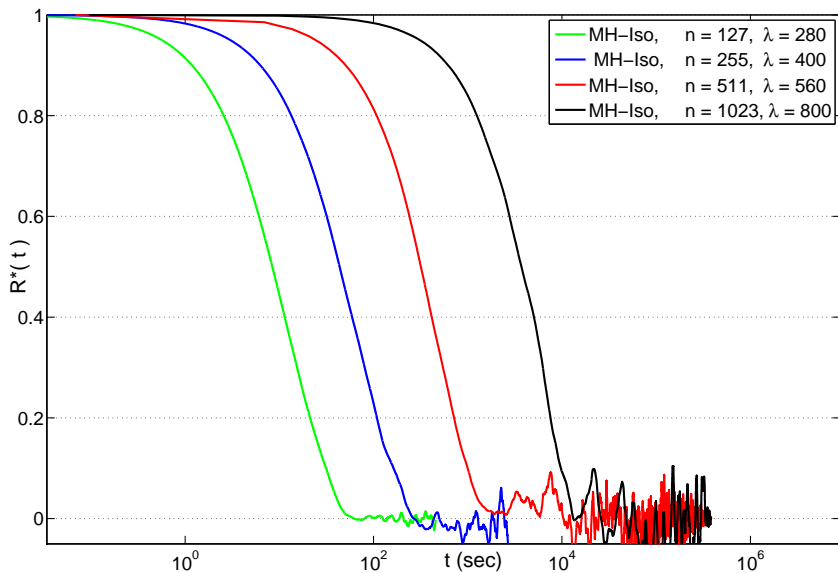


Figure: Temporal autocorrelation plots  $R^*(t)$  for MH Sampler.

## Total Variation Deblurring Example in 1D (from Lassas & Siltanen, 2004)

Results:

- ▶ Efficiency of MH samplers dramatically decreases when  $\lambda$  or  $n$  increase.
- ▶ Even for moderate  $n$ , most inference procedures become infeasible.

What else can we do?

- ▶ More sophisticated variants of MH sampling?
- ▶ Sample surrogate hyperparameter models?
- ▶ Try out the other basic scheme: Gibbs sampling.

## Total Variation Deblurring Example in 1D (from Lassas & Siltanen, 2004)

Results:

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What else can we do?

- ▶ More sophisticated variants of MH sampling?
- ▶ Sample surrogate hyperparameter models?
- ▶ **Try out the other basic scheme: Gibbs sampling.**

## Single Component Gibbs Sampling

Given: Density  $p(x)$ ,  $x \in \mathbb{R}^n$  to sample from.

Let  $x_0 \in \mathbb{R}^n$  be an initial state. Define burn-in size  $K_0$  and sample size  $K$ .

For  $i = 1, \dots, K_0 + K$  do:

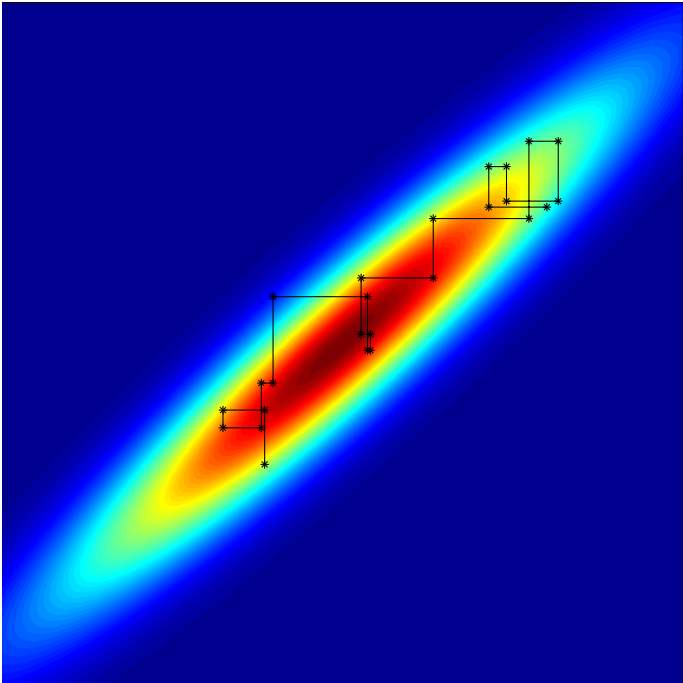
- 1 Set  $x_i := x_{i-1}$ .
- 2 For  $j = 1, \dots, n$  do:
  - (i) Draw  $s$  randomly from  $\{1, \dots, n\}$  (random scan).
  - (ii) Draw  $(x_i)_s$  from the conditional, 1-dim density  $p(\cdot | (x_i)_{[-s]})$ .

Return  $x_{K_0+1}, \dots, x_K$ .

In order to be fast one needs to be able

1. to compute the 1-dim distributions fast and explicit.
2. to sample from 1-dim distributions fast, robust and exact.

Point 2. turned out to be rather **nasty, involved and time consuming to implement**  $\rightsquigarrow$  Details can be found in the paper.



## Total Variation Deblurring Example in 1D (from Lasso & Siltanen, 2004)

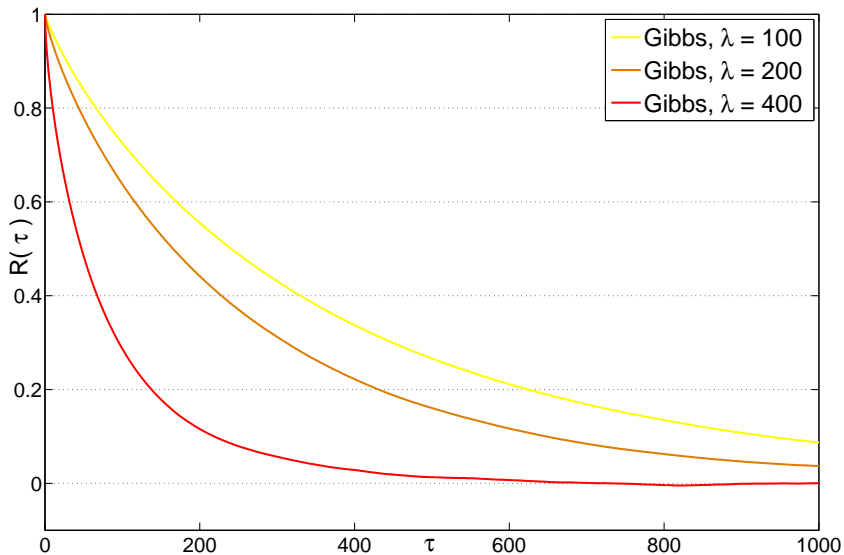


Figure: Autocorrelation plots  $R(\tau)$  for Gibbs Sampler and  $n = 63$ .

## Total Variation Deblurring Example in 1D (from Lassas & Siltanen, 2004)

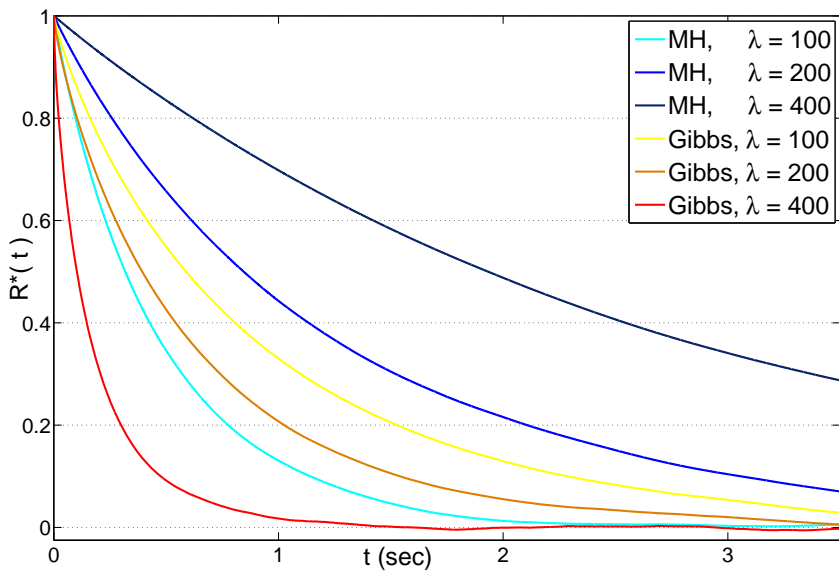


Figure: Temporal autocorrelation plots  $R^*(t)$  for  $n = 63$ .

## Total Variation Deblurring Example in 1D (from Lassas & Siltanen, 2004)

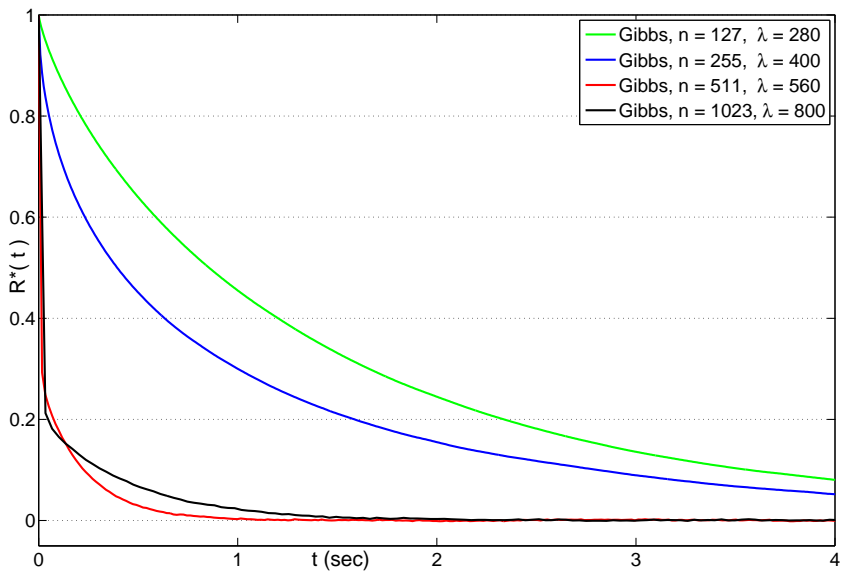


Figure: Temporal autocorrelation plots  $R^*(t)$  for Gibbs Sampler



## Total Variation Deblurring Example in 1D (from Lassas & Siltanen, 2004)

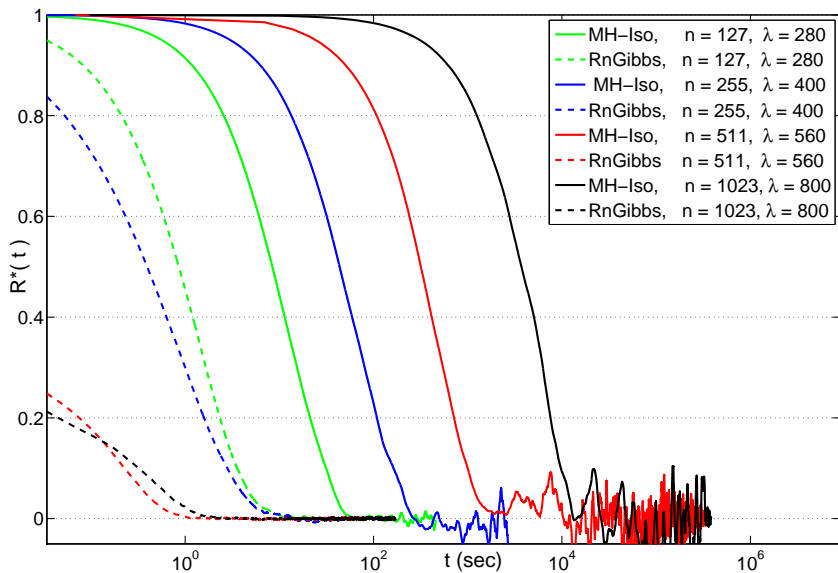
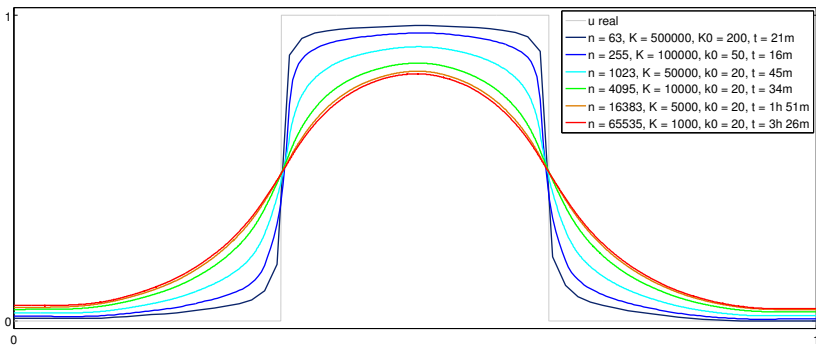


Figure: Temporal autocorrelation plots  $R^*(t)$ .

## Total Variation Deblurring Example in 1D (from Lasso & Siltanen, 2004)

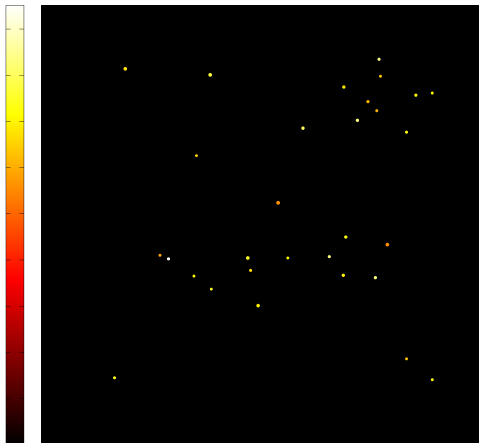
New sampler can be used to address theoretical questions:

- ▶ Lasso & Siltanen, 2004: For  $\lambda_n \propto \sqrt{n+1}$ , the TV prior converges to a smoothness prior in the limit  $n \rightarrow \infty$ .
- ▶ MH sampling to compute CM estimate for  $n = 63, 255, 1023, 4095$ .
- ▶ Even after a month of computation time only partly satisfying results.

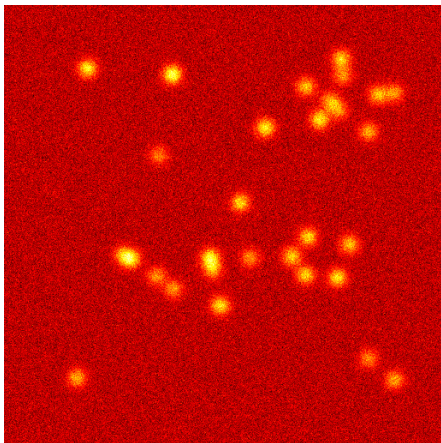


**Figure:** CM estimate computed for  $n = 63, 255, 1023, 4095, 16383, 65535$  using Gibbs sampler on a comparable CPU.

## Image Deblurring Example in 2D



Unknown function  $\tilde{u}$



Measurement data  $m$

- ▶ Gaussian blurring kernel
- ▶ Relative noise level of 10%
- ▶ Reconstruction using  $n = 511 \times 511 = 261\,121$ .

## Image Deblurring Example in 2D

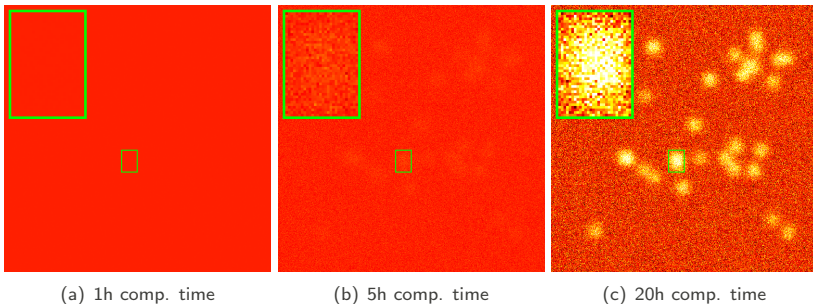


Figure: CM estimates by MH sampler

## Image Deblurring Example in 2D

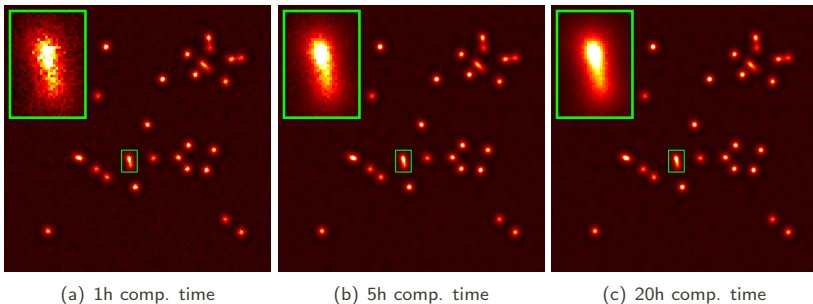


Figure: CM estimates by Gibbs sampler

## Conclusions & Outlook

- ▶ MH is a “black-box sampler”. It may fail dramatically in specific scenarios.
  - ▶ But this is not a general feature of MCMC!
  - ▶ Gibbs sampler incorporate more posterior-specific information into the sampling and perform way better.
  - ▶ Promising results in dimensions larger than any previously reported use for L1-type inverse problems ( $n > 1\,000\,000$  still works...).
- ⇒ Results challenge common beliefs about MCMC in general.

### Work to do:

- ▶ Real applications: Sparse tomography using Besov space priors like in [Kolehmainen, Lassas, Niinimäki, Siltanen, 2012]
- ▶ Tackle theoretical questions, e.g., of how stair-casing in TV can be seen from a Bayesian perspective.
- ▶ Comparison to more sophisticated variants of MH and Gibbs schemes.
- ▶ Generalization to arbitrary  $D$  in  $|Du|_1$ .

Thank you  
for  
your attention!

Full results and all details in:



F. Lucka , 2012.

Fast MCMC sampling for sparse Bayesian inference in high-dimensional inverse problems using L1-type priors

*submitted to Inverse Problems; arXiv:1206.0262v1*

- ▶ More sampling methods.
- ▶ Nasty details of the Gibbs sampler!
- ▶ 2D deblurring with  $n = 511^2 = 261\,121$ .
- ▶ Implementation and code.